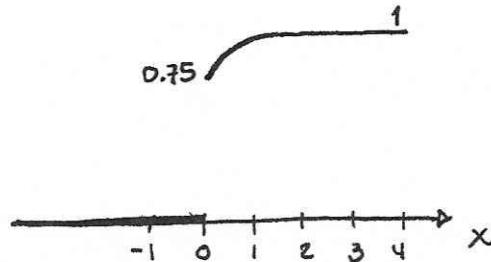


~~4.11~~

~~a)~~



Mixed type random variable

$$(b) P[X \leq 2] = 1 - \frac{1}{4} e^{-2}$$

$$= 0.9662$$

$$P[X=0] = 1 - \frac{1}{4} e^{-0}$$

$$= 0.75$$

$$P[X < 0] = 0$$

$$P[2 < X < 6] = P[X \leq 6] - P[X \leq 2]$$

$$= 1 - \frac{1}{4} e^{-6} - 1 + \frac{1}{4} e^{-2}$$

$$= 0.0332$$

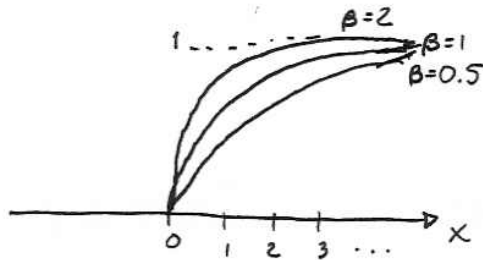
$$P[X > 10] = 1 - P[X \leq 10]$$

$$= 1 - \left(1 - \frac{1}{4} e^{-10} \right)$$

$$= 1.135 \times 10^{-5}$$

4.13

~~a)~~



(b)
$$P[j\lambda < X < (j+1)\lambda] = P[X \leq (j+1)\lambda] - P[X \leq j\lambda]$$

$$= 1 - e^{-\left(\frac{(j+1)\lambda}{\lambda}\right)^\beta} - \left(1 - e^{-\left(\frac{j\lambda}{\lambda}\right)^\beta}\right)$$

$$= e^{-j^\beta} - e^{-(j+1)^\beta}$$

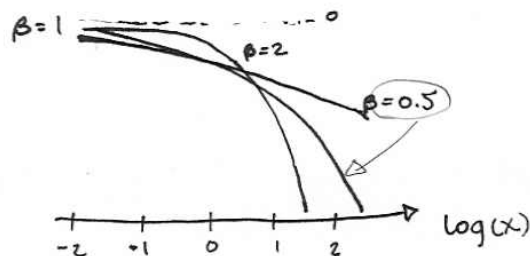
$$= \begin{cases} 0 & j < -1 \\ 1 - e^{-(j+1)^\beta} & -1 \leq j < 0 \\ e^{-j^\beta} - e^{-(j+1)^\beta} & j \geq 0 \end{cases}$$

$$P[X > j\lambda] = 1 - P[X \leq j\lambda] = 1 - \left(1 - e^{-\left(\frac{j\lambda}{\lambda}\right)^\beta}\right)$$

$$= \begin{cases} 0 & j < 0 \\ e^{-j^\beta} & j \geq 0 \end{cases}$$

~~c)~~
$$\log P[X > x] = \begin{cases} \log(1) & x < 0 \\ \log\left(e^{-\left(\frac{x}{\lambda}\right)^\beta}\right) & x \geq 0 \end{cases}$$

$$= \begin{cases} 0 & x < 0 \\ -\left(\frac{x}{\lambda}\right)^\beta & x \geq 0 \end{cases}$$



4.2 The Probability Density Function

4.14 $1 = \int_{-1}^1 f_X(x) dx = c \int_{-1}^1 x(1-x^4) dx = c \left[x - \frac{x^5}{5} \right]_{-1}^1 = \frac{8}{5}c \Rightarrow c = \frac{5}{8}$

b) $F_X(x) = 0$ for $x < -1$; $F_X(x) = 1$ for $x > 1$
 For $-1 \leq x \leq 1$

$$F_X(x) = \int_{-1}^x \frac{5}{8}(1-x'^4) dx' = \frac{1}{2} + \frac{1}{8}(5x - x^5)$$

c) $P[|X| < \frac{1}{2}] = P[-\frac{1}{2} < X < \frac{1}{2}] = F_X(\frac{1}{2}) - F_X(-\frac{1}{2}) = \frac{79}{8(16)}$

4.15 a) We use the fact that the pdf must integrate to one:

$$1 = \int_0^1 f_X(x) dx = c \int_0^1 x(1-x) dx = c \left[\frac{x^2}{2} - \frac{x^3}{3} \right]_0^1 = \frac{c}{6} \Rightarrow c = 6$$

b) $P\left[\frac{1}{2} \leq X \leq \frac{3}{4}\right] = 6 \int_{1/2}^{3/4} x(1-x) dx = 6 \left[\frac{x^2}{2} - \frac{x^3}{3} \right]_{1/2}^{3/4} = 0.34375$

c) For $x < 0$, $F_X(x) = 0$; for $x > 1$, $F_X(x) = 1$
 For $0 \leq x \leq 1$

$$F_X(x) = \int_0^x f_X(x') dx' = 3x^2 - 2x^3$$

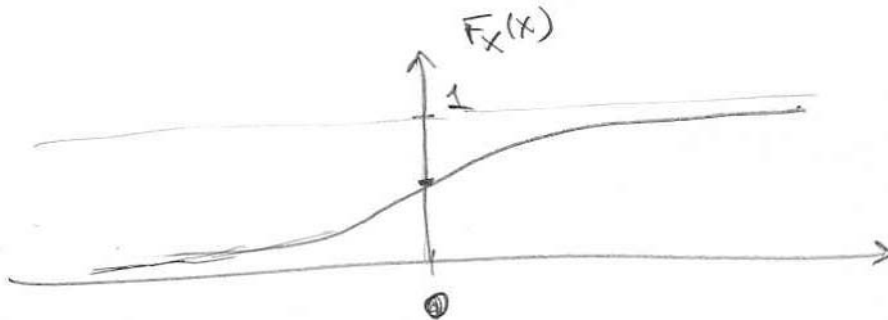
4.23

$$f_X(x) = \frac{\alpha/\pi}{x^2 + \alpha^2}$$

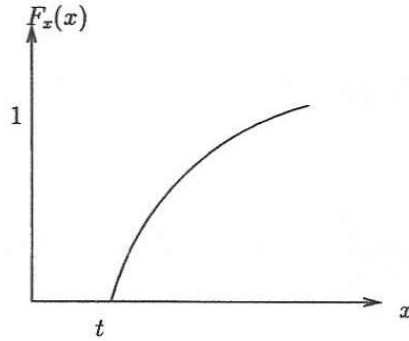
$$F_X(x) = \int_{-\infty}^x \frac{\alpha/\pi}{t^2 + \alpha^2} dt = \int_{-\infty}^x \frac{1/\pi}{1 + (\frac{t}{\alpha})^2} d(\frac{t}{\alpha})$$

$$= \frac{1}{\pi} \tan^{-1}\left(\frac{t}{\alpha}\right) \Big|_{-\infty}^x$$

$$= \frac{1}{\pi} \left[\tan^{-1}\left(\frac{x}{\alpha}\right) + \frac{\pi}{2} \right] \quad -\infty < x < \infty$$



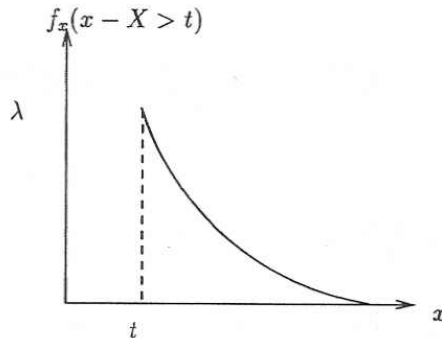
4.29 (a)



$$\begin{aligned}
 F_X(x|X > T) &= \frac{P[\{X \leq x\} \cap \{X > t\}]}{P[X > t]} \\
 &= \begin{cases} 0 & x < t \\ \frac{F_X(x) - F_X(t)}{1 - F_X(t)} & x \geq t \end{cases} \\
 &= \begin{cases} 0 & x < t \\ \frac{(1 - e^{-\lambda x}) - (1 - e^{-\lambda t})}{1 - (1 - e^{-\lambda t})} & x \geq t \end{cases} \\
 &= \begin{cases} 0 & x < t \\ \frac{e^{-\lambda x} - e^{-\lambda t}}{e^{-\lambda t}} & x \geq t \end{cases}
 \end{aligned}$$

$F_X(x|x > t)$ is delayed version of $F_X(x)$

~~(b)~~ $f_X(x|x > t) = \frac{f_X(x)}{1 - F_X(t)} = \frac{\lambda e^{-\lambda x}}{e^{-\lambda t}} = \lambda e^{-\lambda(x-t)}, \quad x \geq t$



(c)

$$\begin{aligned}
 P &= [X > t + x | X > t] \quad x \geq 0 \\
 &= \frac{P[\{X > t + x\} \cap \{X > t\}]}{P[X > t]} \\
 &= \frac{1 - F_X(t + x)}{1 - F_X(t)} \\
 &= \frac{1 - (1 - e^{-\lambda(t+x)})}{1 - (1 - e^{-\lambda t})} \\
 &= e^{-\lambda x} \\
 &= P[X > x]
 \end{aligned}$$

Additional waiting time does not depend on time already spent waiting \Rightarrow "memoryless"

4.49

$$\begin{aligned}
 Y &= A \cos \omega t + c \\
 E[Y] &= E[A] \cos \omega t + c = m \cos \omega t + c \\
 \text{VAR}[Y] &= \cos^2 \omega t \text{VAR}[A] = \sigma^2 \cos^2 \omega t
 \end{aligned}$$

4.50
 3.11
 (a)

$$\begin{aligned}
 \mathcal{E}[Y] &= \int_{-\infty}^{\infty} g(x) f_X(x) dx \quad \text{write integral into three parts} \\
 &= -a \int_{-\infty}^{-a} f_X(x) dx + \int_{-a}^a x f_X(x) dx + a \int_a^{\infty} f_X(x) dx \\
 &= -a F_X(-a) + \int_{-a}^a x f_X(x) dx + a(1 - F_X(a^-)) \\
 \mathcal{E}[Y^2] &= a^2 F_X(-a) + \int_{-a}^a x^2 f_X(x) dx + a^2(1 - F_X(a^-)) \\
 \text{VAR}[Y] &= \mathcal{E}[Y^2] - \mathcal{E}[Y]^2
 \end{aligned}$$

(4.50) (b)

$$\mathcal{E}[Y] = \underbrace{-(-1)P[Y \leq -1]}_{=0} + \underbrace{(1)P[Y \geq 1]}_{=0} + \int_{-1}^1 x \underbrace{\frac{1}{2} e^{-|x|}}_{\substack{\text{odd} \\ \text{even}}} dx = 0$$

$$\begin{aligned}
 \text{VAR}[Y] &= \mathcal{E}[Y^2] = (-1)^2 P[Y \leq -1] + (1)^2 P[Y \geq 1] \\
 &\quad + \int_{-1}^1 x^2 \frac{1}{2} e^{-|x|} dx \\
 &= e^{-1} + 2 \cdot \frac{1}{2} \int_0^1 x^2 e^{-x} dx \\
 &\quad \underbrace{e^{-x} (x^2 + 2x + 2)}_{\text{from appendix}} \Big|_0^1 \\
 &= e^{-1} + 5e^{-1} - 2 = 6e^{-1} - 2
 \end{aligned}$$

4.52

a) $E[Y] = 2E[X] + 2$
 $VAR[Y] = VAR[2X + 2] = VAR[2X] = 4VAR[X]$

b) Laplacian R.V. $E[X] = 0$
 $VAR[X] = \frac{2}{\alpha^2}$

$E[Y] = 2$
 $VAR[Y] = 4\left(\frac{2}{\alpha^2}\right) = \frac{8}{\alpha^2}$

c) Caussian R.V. $E[X] = m$
 $VAR[X] = \sigma^2$

$E[Y] = 2m + 2$
 $VAR[Y] = 4\sigma^2$

d) $E[X] = b \int_0^1 \cos(2\pi u) du = -b \sin(2\pi u) \Big|_0^1 = 0$
 $VAR[X] = b^2 \int_0^1 \cos^2(2\pi u) du$
 $= b^2 \int_0^1 \frac{1}{2} du + \frac{b^2}{2} \int_0^1 \cos 4\pi u du$
 $= b^2 \frac{1}{2} + b^2 \left(\frac{1}{4\pi}\right) (-\sin 4\pi u) \Big|_0^1$
 $= \frac{b^2}{2}$

$E[Y] = 2$
 $VAR[Y] = \frac{4b^2}{2}$

4.53

$E[X^n] = \int_0^1 x^n dx = \frac{x^{n+1}}{n+1} \Big|_0^1 = \frac{1}{n+1}$
 $E[Y^n] = \frac{1}{b-a} \int_a^b y^n dy = \frac{1}{b-a} \left[\frac{b^{n+1} - a^{n+1}}{n+1} \right]$

~~4.65~~ Gamma RV

(a)
$$E[X] = \int_0^{\infty} x \frac{\lambda (\lambda x)^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} dx = \int_0^{\infty} \frac{\lambda (\lambda x)^{\alpha} e^{-\lambda x}}{\lambda \Gamma(\alpha)} dx$$

$$= \int_0^{\infty} \frac{\Gamma(\alpha+1)}{\lambda \Gamma(\alpha)} \frac{\lambda (\lambda x)^{\alpha} e^{-\lambda x}}{\Gamma(\alpha+1)} dx = \frac{\Gamma(\alpha+1) \Gamma(\alpha)}{\lambda} \underbrace{\int_0^{\infty} \frac{\lambda (\lambda x)^{\alpha} e^{-\lambda x}}{\Gamma(\alpha+1)} dx}_1$$

$$= \frac{\alpha}{\lambda}$$

(b)
$$E[X^2] = \int_0^{\infty} x^2 \frac{\lambda (\lambda x)^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} dx = \frac{\Gamma(\alpha+2)}{\Gamma(\alpha)} \frac{1}{\lambda^2} \underbrace{\int_0^{\infty} \frac{\lambda (\lambda x)^{\alpha+1} e^{-\lambda x}}{\Gamma(\alpha+2)} dx}_1$$

$$= \frac{(\alpha+1)\alpha}{\lambda^2}$$

$$\Gamma(\alpha+2) = (\alpha+1)\Gamma(\alpha+1)$$

$$= (\alpha+1)\alpha\Gamma(\alpha)$$

$$\text{VAR}[X] = \frac{(\alpha+1)\alpha}{\lambda^2} - \frac{\alpha^2}{\lambda^2}$$

$$= \frac{\alpha}{\lambda^2}$$

~~(c)~~ m-Erlang $\alpha = m$

$$E[X] = \frac{m}{\lambda} \quad \text{VAR}[X] = \frac{m}{\lambda^2}$$

~~(d)~~ chi-square $\alpha = k/2 \quad \lambda = 1/2$

$$E[X] = \frac{k}{2} \frac{1}{1/2} = k$$

$$\text{VAR}[X] = \frac{k}{2} \frac{1}{1/4} = 2k$$

4.73 $P[Z \leq 3] = P[X^2 \leq 3] = P[X \leq \sqrt{3}]$ since $X \rightarrow$ non-negative

$$f_Z(z) = \frac{d}{dz} F_X(\sqrt{z}) = f_X(\sqrt{z}) \frac{1}{2} z^{-1/2}$$

$$= \frac{f_X(\sqrt{z})}{2\sqrt{z}} = \frac{\sqrt{z}}{2\sqrt{z}} \frac{1}{\alpha^2} e^{-\sqrt{z}^2/2\alpha^2} = \frac{1}{2\alpha^2} e^{-z/2\alpha^2}$$

exponential RV.

4.74

$$P[N=0] = F_X(\pi)$$

$$P[N=1] = F_X(2\pi) - F_X(\pi)$$

$$\vdots$$

$$P[N=n] = F_X((n+1)\pi) - F_X(n\pi)$$

$$= 1 - e^{-\lambda(n+1)\pi} - (1 - e^{-\lambda n\pi})$$

$$= (e^{-\lambda n\pi} - e^{-\lambda(n+1)\pi})$$

$$= e^{-\lambda n\pi} (1 - e^{-\lambda\pi})$$

$$= e^{-\frac{n\pi}{5\pi}} (1 - e^{-\frac{\pi}{5\pi}})$$

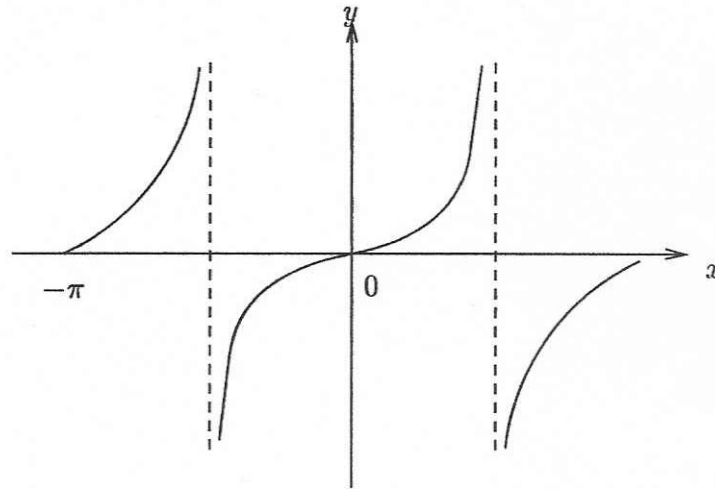
$$= (1 - e^{-1/5}) (e^{-1/5})^n \quad n=0, 1, 2, \dots$$

geometric RV

4.87 $Y = a \tan X.$

$$x = \tan^{-1}(y/a), \quad -\frac{\pi}{2} \leq x \leq \frac{\pi}{2}$$

$$\frac{dx}{dy} = \frac{1}{1 + (y/a)^2} \frac{1}{a} = \frac{a}{y^2 + a^2}$$



$$\begin{aligned} f_X(y) &= \sum_k f_X(x) \left| \frac{dx}{dy} \right|_{x=x_k} \\ &= 2 \cdot \frac{1}{2\pi} \frac{a}{y^2 + a^2} \\ &= \frac{a/\pi}{y^2 + a^2} \end{aligned}$$

Y is a Cauchy RV.

✓
4.88

$$X = -\ln(1-U)$$

$$e^{-X} = (1-U)$$

$$U = 1 - e^{-X}$$

$$\begin{aligned} f_X(x) &= f_U(u) \left| \frac{du}{dx} \right| \\ &= f_U(1 - e^{-x}) |e^{-x}| \\ &= e^{-x} \end{aligned}$$